## **POLYCHROMATIC HAMILTON CYCLES**

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# POLYCHROMATIC HAMILTON CYCLES

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#### Abstract

The edges of the complete graph  $K_n$  are coloured so that no colour appears no more than k times,  $k = \lceil n/A \ln n \rceil$ , for some sufficiently large A. We show that there is always a Hamiltonian cycle in which each edge is a different colour. The proof technique is probabilistic.

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#### 1 Introduction

Let the edges of the complete graph  $K_n$  be coloured so that no edge is coloured more than k = k(n) times. We refer to this as a k-bounded colouring. We say that a Hamilton cycle of  $K_n$  is **polychromatic** if each edge is of a different colour. We say that the colouring is **good** if each edge is of a different colour. Clearly the colouring is good if k = 1 and may not be if k = n - 1, since then we may colour all edges incident with vertex 1 the same colour. The question we address here then is that of how fast can we allow k to grow and still guarantee that a k-bounded colouring is good.

Hahn and Thomassen [2] were the first people to consider this problem and they showed that k could grow as fast as  $n^{1/3}$ . In unpublished work Rödl and Winkler [4] in 1984 improved this to  $n^{1/2}$ . In this paper we make further progress and prove

**Theorem 1** There is an absolute constant A such that if n is sufficiently large and k is at most  $\lfloor n/A \ln n \rfloor$  then any k-bounded colouring is good.

**Proof** Throughout the proof assume that A is a large constant and n is large.

Let

$$B = 10^{1/3} A^{2/3}$$
 and  $D = \frac{4B^2}{A} + 20$ .

Let  $p = \frac{B \ln n}{n}$  and construct a random graph H as follows:

**Step 1:** let  $G = G_{n,p} = ([n], E)$ .

(Recall that  $G_{n,p}$  is the random graph with vertex set  $[n] = \{1, 2, ..., n\}$  in which each possible edge occurs independently with probability p.)

University Libraries Jamegie Mellon University Directorich, PA 15213-3890 Step 2: let Y denote the set of edges whose colour appears more than once in E.

Let H = ([n], E/Y).

Thus no two edges of H are of the same colour. We prove our theorem by showing that

 $\mathbf{Pr}(H \text{ is Hamiltonian }) = 1 - o(1).$ 

(Big O and little o notation refer to  $n \to \infty$ .)

This clearly implies that  $K_n$  must have at least one polychromatic Hamilton cycle, provided n is sufficiently large. The proof can be broken into two lemmas.

For  $v \in [n]$  let  $d_v$  denote the number of edges in Y which are incident with v.

Lemma 1  $\Pr(\exists v \in [n] : d_v \ge D \ln n) = o(1)$ 

**Lemma 2** If starting with  $G = G_{n,p}$  we delete an arbitrary set of edges Y to obtain a graph H and in the process no vertex loses more than  $D \ln n$  edges then H is almost surely Hamiltonian.

Our Theorem is clearly an immediate consequence of these two lemmas.

#### 2 Proof of Lemma 1

Let  $d = d_1$  and let  $S_1, S_2, \ldots, S_m$  be the partition of the edges of  $K_n$  incident with vertex 1 into sets of the same colour  $i = 1, 2, \ldots, m$ . Let  $E_i$  be the set of edges of  $K_n$  which have colour *i*. Let  $|S_i| = l_i$  and  $|E_i| = k_i \leq k$  for i = 1, 2, ..., m.

An edge  $e \in S_i$  is deleted in Step 2 if either

(a) 
$$E \cap S_i = \{e\}$$
 and  $E_i/S_i \neq \emptyset$ 

or

(b)  $e \in E$  and  $|E \cap S_i| \ge 2$ .

Let

 $D_x = \{ \text{ edges incident with vertex 1 which are deleted via case (x)} \},$ 

Observe that if  $i \neq j$  then the sets  $D_x \cap S_i$  and  $D_x \cap S_j$  are independent (as random sets.)

The size of  $D_a$ 

Clearly

$$|D_a \cap S_i| = 0 \text{ or } 1, \quad i = 1, 2, \dots, m.$$

Also

$$\begin{aligned} \mathbf{Pr}(|D_a \cap S_i| = 1) &= l_i p (1-p)^{l_i-1} (1-(1-p)^{k_i-l_i}) \\ &\leq l_i (k_i-l_i) p^2 \\ &\leq (k-1) l_i p^2. \end{aligned}$$

Thus

$$\mathbf{E}(|D_a|) \leq (k-1)p^2 \sum_{i=1}^m l_i$$

$$= (k-1)(n-1)p^{2}$$

$$< \frac{B^{2}\ln n}{A}$$

$$= 10^{2/3}A^{1/3}\ln n.$$

Now by Theorem 1 of Hoeffding [1]

$$\Pr\left(|D_{a}| \geq \frac{2B^{2}\ln n}{A}\right) \leq exp\left\{-\frac{B^{2}\ln n}{3A}\right\}$$
$$\leq n^{-2}.$$

The size of  $D_b$ 

Let  $X_i = |E \cap S_i|$  and  $\delta_i = 1_{X_i \ge 2}$ . Thus

$$|D_b| = \sum_{i=1}^m X_i \delta_i.$$

Now fix  $i \in [m]$ . Unfortunately  $X_i$  and  $\delta_i$  are correlated (positively). So let  $Y_i (= BIN(l_i, p))$  be distributed as  $X_i$  but be independent of it. Then we claim that

 $X_i \delta_i$  is majorised by  $(2 + Y_i) \delta_i$ 

i.e. for all  $u \ge 0$ 

$$\mathbf{Pr}(X_i\delta_i \ge u) \le \mathbf{Pr}((2+Y_i)\delta_i \ge u).$$
(1)

To see this we take 2 independent sequences  $A_1, A_2, \ldots A_l, B_1, B_2, \ldots B_l, l = l_i$ of Bernouilli random variables where each is 1 with probability p and zero with probability 1 - p.

Let

$$\rho = \begin{cases} \min\{r : A_1 + A_2 + \ldots + A_r = 2\} & \text{if } A_1 + A_2 + \ldots + A_l \ge 2\\ \infty & \text{if } A_1 + A_2 + \ldots + A_l \le 1 \end{cases}$$

Let

$$Z_1 = \begin{cases} 2 + B_{\rho+1} + \ldots + B_l & \text{if } \rho < \infty \\ 0 & \text{if } \rho = \infty. \end{cases}$$

 $Z_1$  has the same distribution as  $X_i \delta_i$ .

Let

$$Z_2 = \begin{cases} 2 + B_1 + \ldots + B_l & \text{if } \rho < \infty \\ 0 & \text{if } \rho = \infty. \end{cases}$$

 $Z_2$  has the same distribution as  $(2 + Y_i)\delta_i$  and (1) follows immediately. Thus  $|D_b|$  is majorised by  $\sum_{i=1}^{m} (2 + Y_i)\delta_i$ .

Now

$$\mathbf{Pr}(\delta_i = 1) \le \binom{l_i}{2} p^2$$

and so

$$\begin{split} \mathbf{E}(\sum_{i=1}^{m} \delta_i) &\leq p^2 \sum_{i=1}^{m} \binom{l_i}{2} \\ &\leq p^2 \frac{n}{k} \binom{k}{2} \\ &\leq \frac{B^2}{2A} \ln n. \end{split}$$

Hence

$$\Pr\left(\sum_{i=1}^{n} \delta_i \geq \frac{B^2}{A} \ln n\right) \leq exp\left\{-\frac{B^2}{6A \ln n}\right\}$$
$$\leq n^{-2}.$$

Consider now the distribution of  $\sum_{i=1}^{m} (2 + Y_i) \delta_i$  conditional on  $\sum_{i=1}^{m} \delta_i \leq m_0 = \lfloor (B^2 \ln n) / A \rfloor$ . This is majorised by

$$\frac{2B^2}{A}\ln n + \sum_{i=1}^{m_0} Z_i$$

where  $Z_1, Z_2, \ldots, Z_{m_0}$  are independent binomials BIN(k, p) and so  $Z = \sum_{i=1}^{m_0} Z_i = BIN(m_0k, p)$ . Thus

$$\mathbf{E}(Z) \leq (1+o(1))\frac{B^2}{A}\ln n \frac{n}{A\ln n} \frac{B\ln n}{n}$$
$$= (1+o(1))\frac{B^3}{A^2}\ln n$$
$$\leq 11\ln n$$

So

$$\begin{aligned} \mathbf{Pr}(Z \ge 20 \ln n) &\le & \exp\left\{-\frac{1}{3}\left(\frac{9}{11}\right)^2 11 \ln n\right\} \\ &= & O(n^{-2}). \end{aligned}$$

Hence

$$\Pr\left(d \ge \frac{2B^2}{A}\ln n + \frac{2B^2}{A}\ln n + 20\ln n\right) = O(n^{-2}).$$

Multiplying by a factor n to account for all vertices gives the lemma.

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### 3 Proof of Lemma 2

We modify the proof of Posá [3] to account for the deletion of edges. So assume now that  $G = G_1 \cup G_2 \cup G_3$  where  $G_1$  and  $G_2$  are independent copies of  $G_{n,p/2}$  and where  $G_3$  is an independent copy of  $G_{n,p'}$ , where p' satisfies the equation  $1 - p = (1 - p/2)^2(1 - p')$ .  $G_3$  plays no further role in the analysis. We first show that  $G_1/Y$  almost surely contains a Hamilton path. If it doesn't then there exists  $i \in [n]$  such that

there exists a longest path of  $G_1/Y$  which does not go through i which implies

no longest path of  $\Gamma_i = (G_1/Y)/\{i\}$  has an end-vertex adjacent to i in  $G_1$ . Let this final event be denoted by  $\mathcal{E}_i$ . Then

$$\mathbf{Pr}(G_1/Y \text{ has no Hamilton path }) \leq n \mathbf{Pr}(\mathcal{E}_n).$$
 (2)

Let now P be a longest path of  $\Gamma_n$  and let  $x_0$  be one of its end-vertices. Let END be the set of end-vertices of longest paths of  $\Gamma_n$  which can be obtained from P by a sequence of *rotations* keeping  $x_0$  as a fixed end-vertex. (Given a longest path Q with end-vertices  $x_0, y$  and an edge yv where v is an internal vertex of Q, we obtain a new longest path  $Q' = x_0..vy..w$  where w is the neighbour of v on P between v and y. We say that Q' is obtained from Q by a rotation.)

It follows from Posá [3] that

$$|N(\Gamma_n, END)| < 2|END|, \tag{3}$$

where for a graph  $\Gamma$  and a set  $S \subseteq V(\Gamma)$ 

$$N(\Gamma, S) = \{ w \notin S : \exists v \in S \text{ such that } vw \in E(\Gamma) \}.$$

**CLAIM:** with probability  $1-o(n^{-1})$ 

$$S \subseteq [n-1], |S| \le \frac{n}{4D \ln n}$$
 implies  $|N(G_1/\{n\}, S)| \ge 3D(\ln n)|S|.$ 

(The proof of this claim is deferred to the end of the proof of the lemma.) Hence in  $\Gamma_n$  we have with probability 1-o $(n^{-1})$ 

$$S \subseteq [n-1], |S| \le \frac{n}{4D \ln n}$$
 implies  $|N(\Gamma_n, S)| \ge D(\ln n)|S|.$ 

It follows from (3) that with probability  $1-o(n^{-1})$ 

$$|END| \ge \frac{n}{12}$$

Now consider the edges of  $G_1$  from vertex n to END. They are independent of END and so are distributed as B(|END|, p/2). Thus their expected number is at least  $(B \ln n)/24$ . Thus if A and hence B is large there will be at least  $(B \ln n)/48$  such edges with probability  $1 \cdot o(n^{-1})$ . But for large A, D < B/48 and so not all of these edges can be included in Y. Thus  $Pr(\mathcal{E}_n) = o(n^{-1})$  and (2) implies that  $G_1/Y$  almost surely has a Hamilton path.

To finish the proof take a Hamilton path P of  $G_1$  and fix one of its endvertices,  $x_0$  say, and using rotations create a set of end-vertices END of Hamilton paths with one end-vertex  $x_0$ . The above analysis shows that  $|END| \geq \frac{n}{12}$  almost surely. Now add the edges of  $G_2$ , which are independent of  $x_0$  and END. Again we can argue that there are almost surely too many  $x_0 - END$  edges in  $G_2$  for them all to be included in Y and the lemma follows since the existence of any one not in Y means that H is Hamiltonian.

#### **Proof of CLAIM**

If the condition in the claim does not hold then there exist disjoint sets  $S, T \subseteq [n-1], s = |S| \leq n/(4D \ln n), t = |T| \leq 3D(\ln n)s \leq 3n/4$  such that each vertex of T is adjacent to at least one vertex in S and no vertex in  $[n-1]/(S \cup T)$  is adjacent to any vertex of S.

Fix s,t and let  $t_0 = 3sD(\ln n)$  Then the probability of the above event is bounded by

$$\binom{n-1}{s} \binom{n-1}{t} \left(\frac{sp}{2}\right)^t \left(1 - \frac{p}{2}\right)^{s(n-1-s-t)} \leq \left(\frac{ne}{s}\right)^s \left(\frac{ne}{t}\right)^t \left(\frac{sp}{2}\right)^t e^{-snp/10}$$
$$= \left(\frac{ne}{s}\right)^s \left(\frac{e}{t}\right)^t \left(\frac{sB\ln n}{2}\right)^t n^{-sB/10}$$

$$\leq \left(\frac{ne}{s}\right)^{s} \left(\frac{e}{t_{0}}\right)^{t_{0}} \left(\frac{sB\ln n}{2}\right)^{t_{0}} n^{-sB/10}$$
$$= \left(\frac{ne}{s} n^{3D\ln(Be/6D) - B/10}\right)^{s}$$
$$= o(n^{-3})$$

for large A. Now multiply this upper bound by  $n^2$ , which bounds the number of possible s, t, in order to prove the claim.

Finally, we remark that we believe the following

**Conjecture:** there exists an absolute constant  $\epsilon > 0$  such that if  $k < \epsilon n$  then any k-bounded colouring of  $K_n$  is good.

Hahn and Thomassen made a somewhat stronger conjecture.

## References

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